

A stochastic optimisation algorithm for pairwise likelihood estimation of ordinal factor models

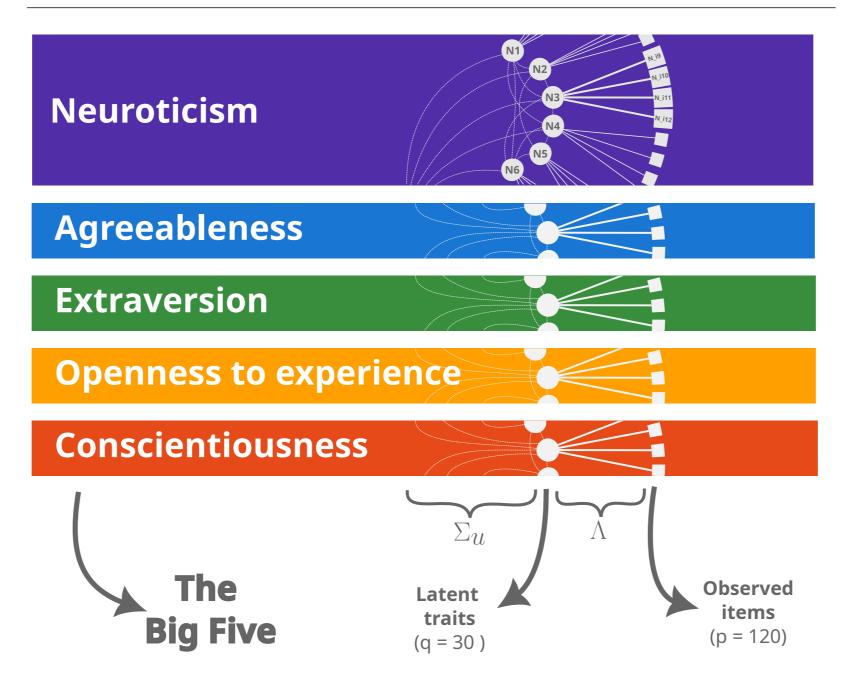
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Summary

Motivated by a big web-based personality test, we develop a stochastic optimisation algorithm that exploits pairwise likelihood structure to scale ordinal factor models to large datasets.

The Big Five Personality test

- Large-scale web-based test designed to measure 5 personality areas [1]: Neuroticism (N), Agreeableness (A), Extraversion (E), Openness to experience (O) and Conscientiousness (C).
- Each area can be further split in 6 personality facets, for a total of 30 mutally correlated latent traits to account for.
- The dataset consists of answers to 120 items on a five-category rating scale, observed on more than 600 thousands units.



Notation

- n units, p observed items, q latent variables;
- Random $n \times p$ matrix of manifest variables Y, with $Y_i = (Y_{i1}, \ldots, Y_{ip}) \in \mathbb{R}^p$ for $i = 1, \ldots, n$. Realizations y, with $y_i = (y_{i1}, \ldots, y_{ip}) \in \mathbb{R}^p$;
- Random $n \times q$ matrix of latent variables U, with $U_i = (U_{i1}, \ldots, U_{iq}) \in \mathbb{R}^q$ for $i = 1, \ldots, n$. Realizations u, with $u_i = (u_{i1}, \ldots, u_{iq}) \in \mathbb{R}^q$ and $u_i \stackrel{iid}{\sim} \mathcal{N}_q(0, \Sigma_u)$;
- Latent covariance matrix $\Sigma_u \in \mathbb{R}^q \times \mathbb{R}^q$. Constrained to be a correlation matrix.
- Loading matrix $\Lambda = (\lambda_1^T, \dots, \lambda_p^T) \in \mathbb{R}^p \times \mathbb{R}^q$ with $\lambda_j = (\lambda_{j1}, \dots, \lambda_{jq}), j = 1, \dots, p.$

Pairwise Likelihood

Stochastic Optimisation

The pairwise log-likelihood **[2]** is constructed exploiting the underlying response variable parameterisation, with:

$$p\ell(\theta) = \sum_{j < j'} \ell_{jj'}(\theta) = \sum_{j < j'} \sum_{s_j, s_{j'}} n_{s_j s_{j'}}^{jj'} \log \pi_{s_j s_{j'}}^{jj'}$$
(1)

where $n_{s_j s_{j'}}^{jj'}$ is the observed frequency of the specific bivariate pattern on columns j, j', while its marginal probability is defined as

$$\pi_{s_j s_{j'}}^{jj'} = \int_{\tau_{s_j-1}}^{\tau_{s_j}} \int_{\tau_{s_{j'}-1}}^{\tau_{s_{j'}}} \phi_2(y_j^*, y_{j'}^*; \rho_{jj'}^{y^*}) dy_{j'}^* dy_j^*; \quad (2)$$

where $\rho_{jj'}^{y^*}$ is the model correlation between variables j and j', computed via $\lambda_j^T \Sigma_u \lambda_{j'}$.

Pros and cons

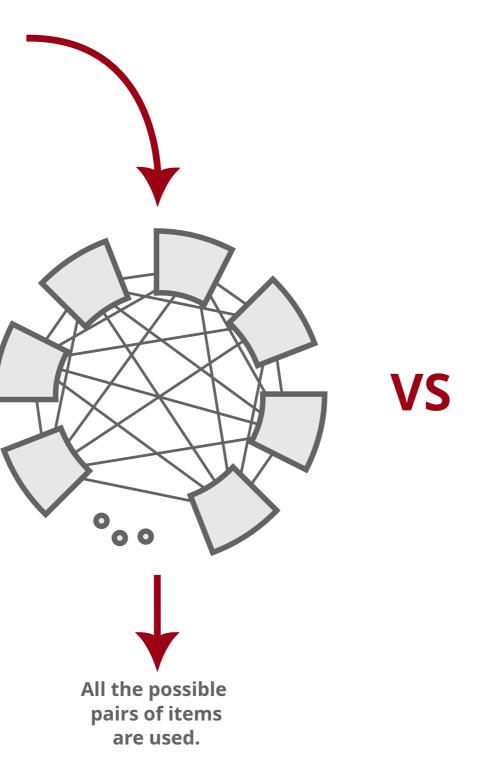
- + It substitutes large-dimensional integration problems with bivariate ones;
- + It reduces data by sufficiency;
- Its computational cost grows with $O(p^2)$;

Details on the Underlying response variable (URV) parameterisation

Assumptions:

- 1. Ordinal responses, $Y_j = s_j \in \{0, ..., c_j 1\}.$
- 2. Data are partial observation of an underlying response normally distributed y^* , such that

$$Y_j = s_j \iff \tau_{s_j-1}^{(j)} < y_j^* < \tau_{s_j}^{(j)},$$





Epoch 1

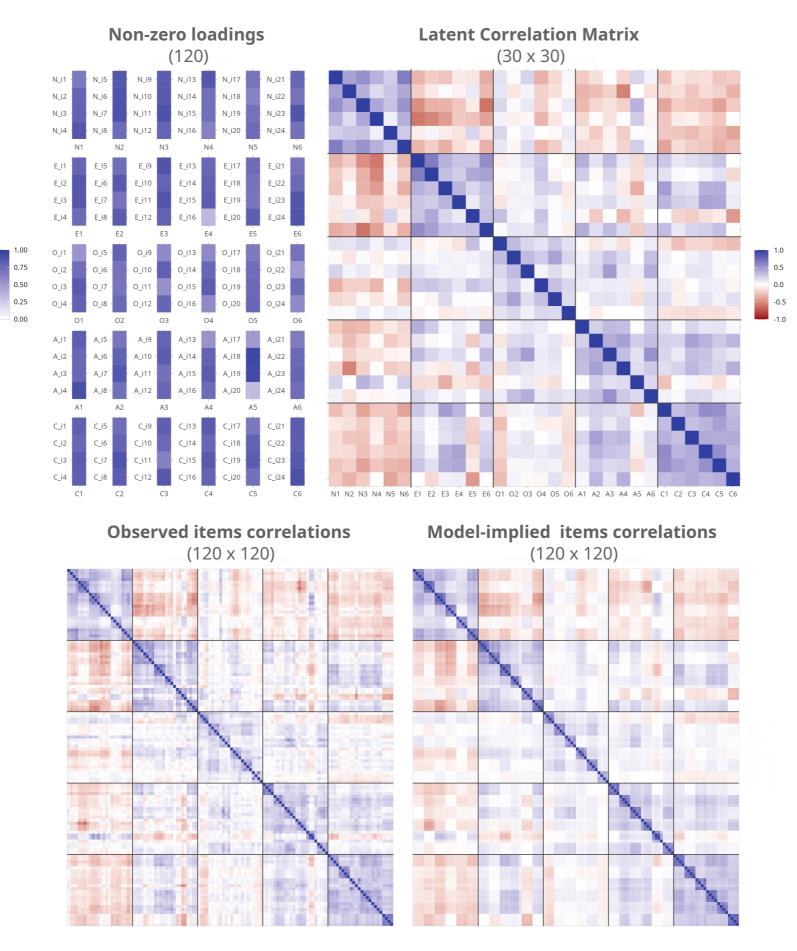
Epoch 3

A different subset

of pairs is drawn

at each epoch.

Epoch 2



Define a stochastic approximation to the gradient via

$$\nabla p\ell(\theta, y) \approx \nabla f(\theta; y, w) = \gamma^{-1} \sum_{j < j'} w_{jj'} \nabla \ell_{jj'}(\theta).$$

- The quantities $w_{jj'}$ are random binary weights such that $w_{jj'} \stackrel{iid}{\sim}$ Bernoulli (γ) .
- Note that, if $\gamma = 1$, we retrieve $\nabla f(\theta; y, w) = \nabla p\ell(\theta)$. If $\gamma \neq 1$ we still have $E_w\{\nabla f(\theta; y, w)\} = \nabla p\ell(\theta)$.

The algorithm

The generic *t*-th epoch alternates [3]:

- 1. Stochastic step: Sample a new set of weights $w^{(t)}$;
- 2. Approximation step: Compute $\nabla f(\theta_{t-1}; y, w^{(t)});$
- 3. Update step: Update θ_t via $\theta_t = \theta_{t-1} + \eta_t \nabla f(\theta_{t-1}; y, w^{(t)})$, where $\eta_t = \eta t^{-.5+\epsilon}$, such that $\sum_{t=1}^{\infty} \eta_t = \infty$ and $\sum_{t=1}^{\infty} \eta_t^2 < \infty$.

At the end of the optimization, trajectories are averaged via $\bar{\theta} = T^{-1} \sum_{t}^{T} \theta_{t}$.

Main Findings

- The proposed framework allows to arbitrarily decrease the complexity per iteration considering only a subset of the pairs. Practitioners can choose the complexity per iteration according to their hardware/time constraints, similarly to what happens when tuning the size of a mini-batch SGD.
- A large-scale factor model application is provided
 with the Big Five Decemplity Test, allowing

and $y^* = \Lambda u + \delta$ with $\Sigma_{\delta} = I_p - \text{diag}(\Lambda \Sigma_u \Lambda^T)$ and $\delta \sim \mathcal{N}_p(0, \Sigma_{\delta})$.

Parameters:

- Loading matrix Λ , latent correlation matrix Σ_u and thresholds vector τ , where $\tau = (\tau^{(1)^T}, \dots, \tau^{(p)^T})^T$ with $\tau^{(j)} = (\tau_0^{(j)}, \dots, \tau_{c_j-2}^{(j)}) \in \mathbb{N}^{c_j-1}$;
- $\theta \in \mathbb{R}^d$ collects parameters from Λ , Σ_u and τ .

Likelihood:

The data marginal likelihood is given by

$$\ell(\theta; y) = \sum_{i=1}^{n} \log \int_{\tau_{s_1-1}^{(1)}}^{\tau_{s_1}^{(1)}} \cdots \int_{\tau_{s_p-1}^{(p)}}^{\tau_{s_p}^{(p)}} \phi_p(y_i^*; \Sigma_{y^*}) dy_i^*.$$
 (3)

where $\phi_p(x; \Sigma)$ is the density of a *p*-dimensional normal distribution evaluated at *x*, with mean zero and variance Σ .

with the Big Five Personality Test, allowing psychometricians to both calibrate items and estimate the latent correlation structure at the same time.

• Future work will focus on:

- Extending the algorithm to proximal updates [3];
- Generalizing to the large class of composite likelihood functions while developing the appropriate inference tools.

References

[1] John, A., Johnson. (2014). Measuring thirty facets of the Five Factor Model with a 120-item public domain inventory: Development of the IPIP-NEO-120, *Journal of Research in Personality*, **51**, 78 – 89.

[2] Katsikatsou, M., Moustaki, I., Yang-Wallentin, F. and Jöreskog, K. G. (2012). Pairwise likelihood estimation for factor analysis models with ordinal data. *Computational Statistics & Data Analysis* **56**, 4243 – 4258.

[3] Zhang, S., Chen, Y. (2022). Computation for Latent Variable Model Estimation: A Unified Stochastic Proximal Framework. *Psychometrika*.

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